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Hans Fehr • Daniela Ujhelyiova

Fertility, Female Labor Supply, and Family Policy

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Fertility, Female Labor Supply, and Family Policy‡

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Abstract

The present paper develops a general equilibrium model with overlapping generations and endogenous fertility in order to analyze the interaction between public policy and household labor supply and fertility decisions.

The model’s benchmark equilibrium reflects the current family policy as well as the differential fertility pattern of educational groups in Germany. Then we simulate alternative reforms of child benefits and family taxation that increase the long-run fertility and growth rate of the economy. Our simulations indicate two central results: First, although households are typically hurt by the first-order effects of family policy, it is possible to generate long-run welfare gains due to positive second-order effects from induced changes in the population structure. Second, specific family policies could be designed that yield a joint increase of the fertility rate and female employment rate as observed in cross-country studies.

JEL Classification: J12, J22

Keywords: stochastic fertility, general equilibrium life cycle model

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1 Introduction

During the past decades most industrialized countries of the Western world have experienced declining fertility rates combined with an increase in female labor market participation (FLMP). Within OECD countries total fertility rates\(^1\) (TFR) decreased from above 2.6 children per woman in 1970 to 1.6 children in 2006, see OECD (2009a). In Germany this negative trend was even stronger. In the old West German states the TFR decreased from its peak of 2.5 children per woman in the mid-1960s to 1.3 children in the mid-1980s (StaBu, 2007). The decline did not turn out so significant in the former German Democratic Republic, so that the TFR has stayed at about 1.4 children per woman since the German reunification. During the last three decades, the mean age of mothers at first childbirth increased, on average from 23.8 to 27.7 years. Although the postponement of first childbirth is very significant with an increase of over one year per decade, cohort fertility data indicates that recuperation is only partial at higher ages. Since women give fewer birth to children and later in life, it seems natural that they increase their share in the labor force at the same time. As documented by OECD (1995, 2009b) participation rates of prime-age women rose significantly within OECD countries from 48.3 percent in 1974 to 61.3 percent in 2008. The inverse relationship between fertility and FLMP can also be observed in the cross-section data with respect to the individual skill level. For example, in 2007 participation rates in Germany increased from 57 percent for low-skilled women up to 85 percent for high-skilled women while at the same time fertility rates declined from 1.9 children per low-skilled woman to 1.1 children per high-skilled woman (StaBu, 2009b,c).

At first sight, both trends could be explained quite well by Becker’s (1965) seminal work on household time allocation. According to this model the rising female earnings power has increased the opportunity cost of child bearing and leisure, which in turn has reduced the demand for both goods. These theoretical results fit in quite well with empirical studies that document the positive correlation between women’s education and their labor market participation (OECD, 2009b), postponed maternity (Gustafsson and Kalwij, 2006) and childlessness (Hoem, Neyer and Andersson, 2006). However, the relationship between women’s fertility and labor supply decisions becomes more complex when examining time series data of OECD countries. First, the trend that women give fewer birth to children and later in life is not uniform among all OECD countries. In countries such as Belgium, France, the Scandinavian countries and the United States the TFR has either remained at or has recently recovered above 1.8 children per woman, while in countries such as Germany, Italy or Spain fertility rates have remained constant below 1.4 children for many years. Second, several recent studies (Ahn and Mira, \(^1\)The average number of children that would be born alive to a woman during her lifetime if she were to pass through her childbearing years conforming to the age-specific fertility rates of a given year.
2002; Del Boca, Pasqua and Pronzato, 2009) have stressed that across many OECD countries the relationship between female employment and fertility has changed over the last 25 years. While in 1980 there was a clear negative correlation between female employment and total fertility rates, in 2005 some OECD countries with higher rates of female employment also had relatively high birth rates, so that the correlation in these countries is now positive. Finally, the relation between education and fertility has changed its sign recently as well. In the past, OECD countries with higher rates of women’s enrollment in tertiary education were also those featuring lower fertility rates. In the 1990s, however, countries with higher women’s education also reported higher fertility rates (d’Addio and d’Ercola, 2005).

Since many industrialized countries are facing an enormous pressure due to the ageing population and the decline in population size, increasing TFRs and FLMPs jointly is extremely important for future labor markets and social security systems. Not surprisingly, family policy is attracting an increasing public attention while at the same time a substantial amount of both theoretical and empirical research aims to uncover the central determinants of a woman’s joint childbearing and labor supply decisions. Extending the overlapping generation model of Galor and Weil (1996), Martinez and Iza (2004) focus on labor market conditions, technological change and private child care provisions in order to explain the interaction between differential birth rates and women’s labor supply in a growing economy. On the other hand, Apps and Rees (2004) discuss differences in the public child care systems and family taxation in order to explain the described heterogeneity in fertility rates and labor market outcomes across countries. Recent empirical studies by Björklund (2006), Del Boca et al. (2009), Kalwij (2009), Laroque and Salanie (2004) and Lalive and Zweimüller (2009) indicate that various policies that reduced the opportunity cost of children were indeed successful in increasing fertility rates and labor market participation for women. Another strand of the literature has focussed on the differential fertility of different educational groups. Since intergenerational educational mobility is fairly low (Woessmann, 2008), de la Croix and Doepke (2003) develop a model in which persistent intragenerational fertility differences deteriorate the income distribution and reduce future economic growth.

The present study is related to the recent literature of calibrated models on the economics of the family. In this context, Caucutt, Guner and Knowles (2002), Greenwood, Guner and Knowles (2003) or Guner and Knowles (2009) distinguish two sexes and analyze the interaction between the marital status, employment, childbearing and human capital investment in a three-stage decision process. In these models women’s fertility declines in their educational level since children are time intensive and thus more costly for women with high productivity. The models are applied to analyze the impact of changes in women’s productivity or of different government policies on marriage, fertility, employment, education and the overall income distribution. Erosa, Fuster and Restuccia (2002) abstract from the marital decision in order to
study the impact of fertility and labor market decisions on human capital accumulation in a search theoretic framework with job mobility and different job qualities. The study finds that fertility decisions, which generate long lasting employment and wage effects, mainly explain the observed gender wage gap in the U.S. while tenure capital plays only a minor role. Da Rocha and Fuster (2006) apply a similar search model but they consider an overlapping generations economy where only adult females make childbirth and labor market decisions. Due to labor market frictions the model is able to generate the observed positive correlation between fertility and employment among OECD countries mentioned above. All studies discussed so far are partial equilibria since they mainly consider the household side, abstract from capital accumulation, public budgets and endogenous factor price repercussions. In contrast, the present study builds on Conesa (2000) or Doepke, Hazan and Maoz (2008) who abstract from labor market search and analyze the household’s fertility and women’s labor supply decisions within a dynamic general equilibrium framework. Conesa (2000) focusses on intragenerational differences in fertility behavior and replicates the delayed childbirth and lower birth rates of higher educated women. Doepke et al. (2008) consider only one representative household per cohort and generate a baby boom by restricting labor market access for women. However, both studies mostly neglect the government sector. This is where the present study steps in. In particular, we analyze reform options for child benefits and family taxation in Germany, since this country has an extremely negative past record regarding TFRs and FLMPs. Due to the fact that these determinants are negatively correlated, it is questionable whether the government is able to increase fertility and female employment rates simultaneously by specific measures.

The next section motivates the quantitative approach considering a simple static model of fertility choice. Section 3 describes the structure of the simulation model. Section 4 explains the calibration and simulation design. Finally, section 5 presents the numerical results and section 6 offers some concluding remarks.

2 The static model of fertility choice

In order to discuss some central mechanisms that motivate our quantitative approach, this section introduces the most basic model of fertility choice in which children provide direct utility benefits. Households maximize utility subject to a budget constraint where women divide their time endowment (normalized to unity) into working and child-rearing. Assuming

\footnote{See Jones, Schoonbroodt and Tertilt (2008) for a discussion of such static fertility models. Alternatively, children may also be viewed as an investment providing old-age security, see Boldrin and Jones (2002).}
separable utilities and considering the CES case we receive

$$\max_{c,n} U(c, n) = \alpha_c \frac{c^{1-\frac{1}{\rho}}}{1-\frac{1}{\rho}} + \alpha_n \frac{n^{1-\frac{1}{\rho}}}{1-\frac{1}{\rho}} \quad \text{s.t.} \quad c + b_0 n = w(1 - b_1 n)$$

where $c$ and $n$ denote consumption and the number of children, respectively. The parameters $\alpha_c$ and $\alpha_n$ as well as the substitution elasticity $\rho$ determine the preference structure, $w$ defines the wage and $b_0, b_1$ the costs of children in terms of money and time. Note that family policy intends to reduce the costs of children either via $b_0$ (for example by direct transfers per child) or indirectly via $b_1$ (for example by providing subsidized child care). The explicit solution for the number of children is given by

$$n = \frac{w}{\left[\frac{\alpha_c}{\alpha_n}(b_0 + wb_1)\right]^\rho + b_0 + wb_1}.$$

As it turns out, a clearly negative relationship between income and fertility could be generated by ignoring monetary costs of children ($b_0 = 0$) and setting $\rho > 1$. If only time costs of children matter then high-wage families face higher opportunity costs of having children. With a high elasticity of substitution between children and consumption, the substitution effect dominates so that $n(w)$ decreases as in the data. It is also obvious that family policy – that reduces $b_0$ and/or $b_1$ – is able to increase the fertility rate.

However, one has to keep in mind that financing of family policy reduces net income which may either counteract or strengthen the effects on fertility of the policy instruments. In addition, changes in the cost parameters may affect fertility decisions of various income classes differently. While direct payments per child may have a strong income effect on low-skilled households, they have only a negligible impact on high-skilled parents. Quite the opposite applies to family policy instruments that reduce the time costs of children. The simple static model also neglects leisure demand and the interaction between the fertility choice and female labor supply decision.

Although the utility from leisure consumption could also be considered in the static model, the labor supply decision in the present context has to account for the timing of births and the accumulation and depreciation of female human capital during child rearing. Consequently, the joint fertility and labor supply decision has to be analyzed in a dynamic framework. In such a setup it is also possible to quantify the macroeconomic growth effects resulting from the adjusted fertility pattern. The next section discusses the structure of such a simulation model.
3 The dynamic model economy

3.1 Demographics and intracohort heterogeneity

We consider an economy populated by overlapping generations of married couples who live for \( J \) periods indexed by adult ages \( j \in J = \{1, \ldots, J\} \). The life cycle of a representative household is described in Figure 1. We assume that both adult members of the household belong to the same skill level \( s \in S = \{1, \ldots, S\} \). Men work continuously until age \( J_R - 1 \), afterwards they retire. All woman retire at the same age as men, but they can choose in every period before retirement how much they work. Apart from the labor supply and savings decision, couples face the decision about the number and timing of their children. Women can give birth to children until age \( J_F \). We abstract from twins, triplets etc., so that only one child per period can be born. Consequently, the total number of children of the age-\( j \) household is \( n_j \in \mathcal{N} = \{0, \ldots, J_F\} \). Parents raise their children for \( J_K \) periods, so that \( k_i \in J \) indicates the age of the \( i \)-indexed child of the household. After birth, all children of a cohort are identical until they reach adulthood.

Figure 1: The child-household life cycle

Our model only considers the long-run equilibrium so that we can omit a time index for all variables. It is solved recursively and an age-\( j \) household faces the state vector

\[
z_j = (s, a_j, D_j),
\]

where \( a_j \in \mathcal{A} = [0, \infty) \) define the household’s assets held at the beginning of age \( j \). The vector \( D_j = (k_1, \ldots, k_{n_j}) \in \mathcal{N}^j \) with \( 0 \leq k_1 \leq k_{n_j} \) contains the demographic characteristics of all children of the household. More specifically, \( k_1 \) and \( k_{n_j} \) denote the age of the youngest and the oldest child of the age-\( j \) household, respectively. Given \( D_j \) we can compute the number of children currently living in the household \( m_j \leq n_j \) (i.e. those children where \( k_i \leq J_K \)) or the number of children with ages equal to or less than 6 years \( (m_{6j}) \). Finally, \( m_{1j} \in \{0, 1\} \) indicates whether the age-\( j \) household currently has a newborn child or not.

Each age-\( j \) cohort is fragmented into subgroups \( \xi(z_j) \), according to the initial distribution (i.e. at \( j = 1 \)), the fertility process and optimal individual decisions. Let \( X(z_j) \) be the corresponding
cumulated measure to $\xi(z_j)$. Hence,

$$\int_{\mathcal{C}} dX(z_j) = 1, \quad \text{for all } j \in \mathcal{J}$$

must hold, as $\xi(z_j)$ is not affected by cohort sizes but only gives densities within cohorts. For the sake of simplification, we define $\mathcal{C} = \mathcal{S} \times \mathcal{A} \times \mathcal{N}^J$ as the set of states. Let $N_j$ denote the number of (2-person-)households in the age-$j$ cohort, then

$$M = \sum_j N_j \int_{\mathcal{C}} m_j(z_j) dX(z_j)$$

measures the aggregate number of children living in households while the endogenous (native) population growth rate $\eta$ can be computed from

$$(1 + \eta)^J = \frac{1}{2} \sum_j N_j \int_{\mathcal{C}} m_{1j}(z_j) dX(z_j)$$

where we have normalized the number of the youngest households to unity, i.e. $N_1 = 1$. In order to have zero or positive population growth, we add an exogenous population growth rate from immigration $\bar{\eta}$ so that in equilibrium cohort numbers can be computed from

$$N_j = (1 + \eta + \bar{\eta})N_{j-1}.$$ \hfill (4)

In the following, we will omit the state indices $z_j$ for every variable whenever possible. Agents are then only distinguished according to their age $j$.

### 3.2 The households’ problem

Since we consider the steady state, households maximize utility at the initial age choosing a contingent plan for consumption, labor supply, the number of children and the timing of births. Following Conesa (2000) the fertility decision is modeled so that every period during the fertile years the household decides whether to have an additional child next period or not. Conditional on having decided to have a newborn next period, this event will only happen with a certain probability. Our model assumes a preference structure that is represented by a time-separable, nested CES utility function. Consequently, the problem can be written recursively so that the household at age $j$ and state $z_j$ solves

$$V(z_j) = \max_{c_j, \ell_j, \hat{n}_j} \left\{ u(c_j, \ell_j, \hat{n}_j) + \beta E \left[ V(z_{j+1}|z_j)^{1-\frac{1}{\gamma}} \right] \right\}^{\frac{1}{1-\frac{1}{\gamma}}}$$ \hfill (5)

by choosing per capita consumption of goods $c_j$, leisure consumption of the mother $\ell_j$ and the family size $\hat{n}_j$. With $n_j$ being the number of children, we define $\hat{n}_j = 2 + n_j$. Expected utility
in future periods is discounted by $\beta$ and the intertemporal elasticity of substitution is defined by $\gamma$. The expectation operator $E$ in (5) indicates that future utilities are computed over the distribution of $D_{j+1}$.

Households maximize (5) subject to the budget constraint (6),

$$a_{j+1} = (1 + r)a_j + (1 - \tau)w_j + m_j b^c - m_{aj} p_j l^f_j + p_j - T(y_j) - (1 + \tau_c) f(m_j) c_j$$

with $a_1 = a_{j+1} = 0$. The constraint (6) reflects how children effect resources of the family in the model. In addition to interest income from savings $r a_j$, households receive gross labor income

$$w_j = w e_j(l^h + e^f c^l f_j)$$

with $e^f = \begin{cases} 1 & n_j = m_j = 0 \\ (1 - \delta_c) e^f_{j-1} & n_j, m_j > 0 \\ e^f_{j-1} & n_j > 0, m_j = 0 \end{cases}$

during their working periods. Given the wage rate for effective labor $w$ and $e_j$ as the age-$j$ productivity, labor supply of the husband $l^h$ is exogenously predetermined$^3$ while the working time of the mother $l^f_j$ is endogenously chosen given the husband’s income, and the number of children in the household. We assume that children in the household reduce productivity of the mother where $\delta_c$ measures the depreciation rate which determines the depreciation factor $e^f_j$. The time endowment of the mother is normalized to one and divided between working, childcare and leisure consumption.

$$1 = l^f_j + \Psi(D_j) + \ell_j$$

The time required for childcare measured by the function $\Psi(\cdot)$ depends on the age structure of children since we assume – following Da Rocha and Fuster (2006) or Doepke et al. (2008) – that younger children are more time intensive than older children. Depending on the number of children, households may also receive direct monetary support $b^c$ such as child benefits or parental leave benefits per child. In addition, they have to pay a fee of $p c$ per (younger than six year) child for external child care during the time the mother is working. Households have to pay social security contributions at a rate $\tau$ on gross family income, income taxes that depend on (family-size related) taxable income $y_j$, and receive public pensions $p_j$ during retirement. Finally, the price of consumption goods $c_j$ includes consumption taxes $\tau_c$ and the total consumption of the household is given by multiplying per capita consumption with $f(m_j) = 1.7 + 0.5 m_j$, see Conesa (2000).

$^3$We do not consider this as a strong assumption given the large body of empirical evidence suggesting very limited reaction of men’s labor supply to tax changes, see Heckman (1993) or Eissa and Hoynes (2004).
3.3 Instantaneous utility and the decision to have children

Similar as Conesa (2000) or Doepke et al. (2008) we define the period utility function by

\[
u(c_j, \ell_j, \hat{n}_j) = \left\{ [c_j^{\alpha_1} (\ell_j - \alpha_1)]^{\frac{1-\frac{1}{\rho}}{1-\frac{1}{\rho}}} + \alpha_2 \hat{n}_j^{1-\frac{1}{\rho}} \right\}^{\frac{1}{1-\frac{1}{\rho}}}. \tag{8}
\]

where \(\alpha_1\) denotes the coefficient of consumption in the sub-utility function and \(\rho\) defines the intratemporal elasticity of substitution between consumption and leisure on the one hand and family size on the other hand, while \(\alpha_2\) defines the age-independent preference parameter for family size.

The fertility decision is modeled similar to the college choice in Heckman, Lochner and Taber (1998). In every period during fertile years \(j < J_F\) each household has to decide whether to have one additional child next period or not. The welfare change of having an additional child or not for household \(z_j\) measured by the equivalent variation can be written as

\[
\frac{V(z_j^1)}{V(z_j^0)} - 1 - \epsilon_z
\]

where \(V(z_j^1)\) and \(V(z_j^0)\) measure utilities from having an additional child or not while additional non-pecuniary (i.e. psychological) gains or cost from additional children, which are not observed by the model, are captured by \(\epsilon_z \sim N(0, \sigma^2)\). We assume that the latter are normally distributed within each skill class with mean zero and variance \(\sigma^2\). Due to the law of large numbers, we can now compute the fraction of households that decide to have an additional child from

\[
P \left( \left\{ \frac{V(z_j^1)}{V(z_j^0)} - 1 - \epsilon_z \right\} \right) = \Phi_{0,\sigma^2} \left[ \frac{V(z_j^1)}{V(z_j^0)} - 1 \right],
\]

where \(\Phi_{0,\sigma^2}\) defines the cumulative normal distribution function with mean zero and variance \(\sigma^2\). Conditional on having decided to have a newborn baby next period, this event will happen with probability \(0 \leq \pi \leq 1\). Following Conesa (2000), we assume that fertility uncertainty is independent of age and skill classes during fertile ages. Note that there is no uncertainty if the household decides not to have a newborn.

3.4 The production side

Firms in this economy use capital and labor to produce a single good according to a Cobb-Douglas production technology \(Y = \theta K^\varepsilon L^{1-\varepsilon}\) where \(Y, K\) and \(L\) are aggregate output, capital and labor, respectively, \(\varepsilon\) is capital’s share in production and \(\theta\) defines a technology parameter. Capital depreciates at a rate \(\delta_k\). Firms maximize profits renting capital and hiring labor from
households such that net marginal products equal $r$ the interest rate for capital and $w$ the wage rate for effective labor.

### 3.5 The government sector

Our model distinguishes between the general government budget and the pension system. In each period of the long-run equilibrium, the government issues new debt $(\eta + \bar{\eta})B_G$ and collects taxes from households in order to finance general government expenditure $G$ (which is fixed per capita), in-kind benefits or services for families $G_c$ and direct monetary support $Mb^c$ to families as well as interest payments on its debt, i.e.

$$ (\eta + \bar{\eta})B_G + T = G + G_c + Mb^c + rB_G, \quad (9) $$

where $T$ defines tax revenues from income and consumption taxation

$$ T = \tau C + \sum_j N_j \int C T(y_j) dX(z_j) \quad (10) $$

with $C$ as aggregate consumption. We assume that contributions to public pensions are exempted from income tax while benefits are fully taxed. Consequently, taxable income $y_j$ is computed from gross labor income net of pension contributions, capital income and – after retirement – public pensions, i.e.

$$ y_j = (1 - \tau) w_j + ra_j + p_j. \quad (11) $$

Given taxable income, we apply the German progressive tax code of the year 2005. In-kind benefits for families are modeled as a fixed cost per child $\kappa$ that covers childcare institutions and the provision of schools and universities minus payments of parents for public childcare, i.e.

$$ G_c = \kappa M - p_c \sum_j N_j \int C m_{\theta j}(z_j) t_{j}(z_j) dX(z_j). \quad (12) $$

In each period, we assume a fixed debt to output ratio and balance the public budget by adjusting the consumption tax rate.

Finally, the pension system pays old-age benefits and collects payroll contributions from wage income. Pension benefits $p_j$ of a retiree household at age $j \geq J_R$ in a specific year are uniform across age and skill classes and computed as a fixed replacement rate of average income $\bar{w}$. Since the budget of the pension system must be balanced in every period by payroll contributions, we have

$$ \tau w L = \sum_{j=J_R}^{J} N_j \int C dX(z_j), \quad (13) $$

where $L$ denotes total labor supply defined in (15) below.
3.6 Equilibrium conditions

Our initial long-run equilibrium is computed in a closed economy so that factor prices are endogenous and the trade balance is zero. Then we implement a policy reform and compute the resulting long-run equilibrium in a small open economy where we keep the factor prices of the initial steady state constant.

In addition to factor prices being equal to marginal products, we need households to maximize (5) with respect to the respective constraints (6) and (7), an invariant measure of households \( \xi \) over the whole state space and market clearance for the capital, labor and goods market in the closed or small open economy:

\[
K + B_G + B_F = \sum_j N_j \int_C a_j(z_j)dX(z_j) \tag{14}
\]

\[
L = \sum_j N_j \int_C \left[e_j^{th} + e_j^{lf}(z_j)\right] dX(z_j) \tag{15}
\]

\[
Y = C + G + (\bar{\eta} + \bar{\eta} + \delta)K + TB, \tag{16}
\]

where \( B_F \) measures foreign debt and \( TB \) denotes the trade balance.

4 Calibration of the initial equilibrium

4.1 Parameterizing the model

Table 1 reports the central parameters of the model. In order to reduce computational time, each model period covers two years. Therefore, children are at home until age 19 (\( J_K = 10 \)), then they start adult life at age 20 (\( j = 1 \)), women can have children until age 36 (\( J_F = 8 \)), households are forced to retire at age 66 (\( J_R = 23 \)) and face a life span of 80 years (\( J = 30 \)).

Since we adjust in our initial equilibrium the exogenous growth rate of households in order to have zero growth (i.e. \( \bar{\eta} = -\eta \)), this cohort structure yields a quite realistic dependency ratio between pensioners and working cohorts of 36.4 percent.

We distinguish \( S = 3 \) educational classes and assume that households only marry within the same skill-class. Based on data estimated from German Socio-Economic Panel (SOEP) of the years 1995-2007 we assume that 25, 55 and 20 percent of the cohort are low-, middle- and high-skilled, respectively. These shares also fit in quite well with the shares reported in StaBu (2009a, 26). SOEP data is also used to compute the efficiency profiles \( e_j \) for skill classes across
Table 1: Parameter selection

<table>
<thead>
<tr>
<th>Demographic parameters</th>
<th>Preference parameters</th>
<th>Technology/Budget parameters</th>
<th>Government parameters</th>
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</thead>
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<td>$J = 30$</td>
<td>$\gamma = 0.5$</td>
<td>$\theta = 1.17$</td>
<td>$B_G/Y = 0.6$</td>
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<td>$\varepsilon = 0.3$</td>
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<td>$\kappa = 0.065\bar{w}$</td>
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<tr>
<td>$S = 3$</td>
<td>$\beta = 1.0$</td>
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<td>$p = 0.55\bar{w}$</td>
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<tr>
<td></td>
<td></td>
<td>$e_j$ see text</td>
<td></td>
</tr>
</tbody>
</table>

the life-cycle.\(^4\) Finally, following Conesa (2000) we assume that 80 percent of those households that wish to have a child will receive one.

With respect to the preference parameters, we set the intertemporal elasticity of substitution $\gamma$ as well as the consumption preference in the Cobb-Douglas sub-utility function $\alpha_1$ to 0.5. The chosen value for the intertemporal elasticity of substitution is within the range of commonly used estimates, see the discussion in İmrohoroğlu and Kitao (2009, p.871). The consumption preference parameter yields quite realistic female labor force participation rates, see Table 3 below. The fertility choice parameters $\rho$ and $\alpha_2$ are calibrated such that the model is consistent with completed fertility and timing of fertility as observed in the data. As explained above, the higher the intratemporal elasticity of substitution between ordinary utility (from goods and leisure consumption) and family size $\rho$, the larger is the difference in completed fertility between the high-skilled and the low-skilled class. The preference for family size $\alpha_2$ determines the level of completed fertility. We have set $\rho$ at 0.65 and fixed $\alpha_2$ at 0.35 that yields both the negative relationship between income and children and the target level for the total fertility rate in the initial equilibrium. Finally, in order to calibrate a realistic capital to output ratio, the discount factor is set at 1.0.

With respect to technology parameters we specify the general factor productivity $\theta = 1.17$ in order to normalize labor income and set the capital share in production $\varepsilon$ at 0.3. The annual depreciation rate for capital is set at 5.9 percent which yields a periodic depreciation rate of $\delta_k = 0.122$. Husbands are assumed to work 40 percent of their time endowment which is typically assumed in quantitative studies, see Auerbach and Kotlikoff (1987). The depreciation of women’s productivity depends on the skill level. We assume a one percent depreciation for low-skill mothers and a two percent depreciation for middle- and high-skill mothers. These

\(^4\)The SOEP data base is described in Wagner, Frick and Schupp (2007). See Fehr, Kallweit und Kindermann (2009) for detailed explanations of the estimation procedure and results.
figures are somewhere between the depreciation rates applied by Da Rocha and Fuster (2006) and Doepke et al. (2008). Finally, we assume that the time costs $\Psi(D_j)$ decrease linearly with age of children. Every mother spends 25 percent of the time endowment with a newborn, 8 percent with every child below school age and 5 percent with every remaining child.

With respect to the government sector we assume a debt-to-output ratio of 60 percent and that taxation of gross income (from labor, capital and pensions) is close to the current German income tax code and the marginal tax rate schedule $T05$ which was introduced in 2005. Consequently, given taxable income $y_j$ of the household and applying the income splitting method, the marginal tax rate rises linearly after the basic allowance of 16,600 € from 15 percent to a maximum of 42 percent when $y_j$ passes 104,000 €. In addition, we also account for the solidarity surcharge of 5.2 percent so that we get

$$T(y_j) = 1.052 \times 2 \times T05(y_j/2).$$

In the initial equilibrium we set $p_c = 0.2$ which yields a realistic revenue of private fees to public childcare and is also close to the figures used by Da Rocha and Fuster (2006). In addition, we assume that 6.5 percent of average income is transferred as child benefits. This figure is quite realistic for Germany, where roughly 2,000 € is payed per child per annum and average income amounts to roughly 30,000 €. The same figure is also assumed for in-kind benefits per capita. Finally, we assume that pensions of each household amount to 55 percent of average wages and fix the per capita costs of general public consumption ($G$) in order to get realistic figures in our benchmark.

### 4.2 The initial equilibrium

Table 2 and Figure 2 report some central indices of the calibrated benchmark equilibrium and the respective figures for Germany in 2007/2008. The upper part of Table 2 shows that the model’s total fertility rate and mean age at childbirth match the German situation quite well. However, first child birth is too late in the model compared to reality. Figure 2 compares the actual and the model’s distribution of family sizes at age 36-38 (i.e. when childbirth is completed). Our model replicates the fact that about 50 percent of German families have either one or no child. However, the fraction of families without children is higher in reality compared to the model. Families with more than one child are captured by the model quite well. With respect to the different skill classes our model reflects the fact that fertility rates decrease with income (i.e. skill level). As one can see in Figure 2 the shares of childless families increase and the shares of families with three or more children decrease with the skill level. As shown by Table 2, the skill-specific fertility rates are realistic.
Table 2: The initial equilibrium

<table>
<thead>
<tr>
<th>Calibration targets</th>
<th>Model solution</th>
<th>Germany 2007/2008 &lt;sup&gt;a&lt;/sup&gt;</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total fertility rate (TFR)</td>
<td>1.46</td>
<td>1.38 &lt;sup&gt;b&lt;/sup&gt;</td>
</tr>
<tr>
<td>Total mean age at childbirth (in years)</td>
<td>29.9</td>
<td>29.8 &lt;sup&gt;b&lt;/sup&gt;</td>
</tr>
<tr>
<td>Total mean age at first child (in years)</td>
<td>27.9</td>
<td>26.1 &lt;sup&gt;c&lt;/sup&gt;</td>
</tr>
<tr>
<td>Skill-specific fertility rates</td>
<td>1.96/1.34/1.11</td>
<td>1.94/1.35/1.14 &lt;sup&gt;c&lt;/sup&gt;</td>
</tr>
<tr>
<td>Skill-specific share of childless (in %)</td>
<td>8/21/27</td>
<td>11/16/26 &lt;sup&gt;c&lt;/sup&gt;</td>
</tr>
<tr>
<td>Government indicators (% of GDP)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>In-kind benefits and services</td>
<td>3.0</td>
<td>3.0 &lt;sup&gt;d&lt;/sup&gt;</td>
</tr>
<tr>
<td>Direct monetary support</td>
<td>3.0</td>
<td>3.0 &lt;sup&gt;d&lt;/sup&gt;</td>
</tr>
<tr>
<td>General government expenditure</td>
<td>15.1</td>
<td>15.0</td>
</tr>
<tr>
<td>Interest payments</td>
<td>3.1</td>
<td>2.7</td>
</tr>
<tr>
<td>Tax revenues</td>
<td>24.2</td>
<td>22.5</td>
</tr>
<tr>
<td>Pension benefits</td>
<td>13.5</td>
<td>11.5</td>
</tr>
<tr>
<td>Pension contribution rate (in %)</td>
<td>19.9</td>
<td>19.9</td>
</tr>
<tr>
<td>Consumption tax rate (in %)</td>
<td>19.0</td>
<td></td>
</tr>
<tr>
<td>Other benchmark coefficients</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Skill-specific mean age at first child (in years)</td>
<td>27.4/28.2/27.9</td>
<td>–</td>
</tr>
<tr>
<td>Capital-output ratio</td>
<td>2.8</td>
<td>3.3</td>
</tr>
<tr>
<td>Interest rate p.a. (in %)</td>
<td>5.2</td>
<td>–</td>
</tr>
<tr>
<td>(Native) Population growth rate p.a. (in %)</td>
<td>-1.5</td>
<td>–</td>
</tr>
</tbody>
</table>

Source: <sup>a</sup>IdW (2009), <sup>b</sup>StaBu(2009a), <sup>c</sup>StaBu (2009b), <sup>d</sup>Rosenschon (2006).

With respect to the calibration of family policy measures we follow the comprehensive study of Rosenschon (2006), where overall public expenditures for family in Germany accumulate to 10.7 percent of GDP. However, many family transfer instruments, which are listed there, are not taken into account by the model. With respect to in-kind transfers such as public childcare services and schools, Rosenschon (2006) reports a figure of roughly 3 percent of GDP. In addition, direct transfers to families including child benefits and parental leave benefits also add up to roughly 3 percent of GDP. The remaining figures are calibrated in order to arrive at a realistic government tax structure and macroeconomic situation. We fix public debt to 60 percent of GDP, so that annual interest payments amount to 3.1 percent of GDP. Since we abstract from growth and deficit financing, tax revenues add up to 24.2 percent of GDP. Private consumption amounts to 64.7 percent of GDP and the endogenous consumption tax rate is 19 percent, so that consumption tax revenues are slightly higher than income tax revenues. The average and marginal tax rate of the latter across the total population are 8.9 and 22.6 percent, respectively. However, skill-specific average income tax rates increase from 2.1 to 9.1 and 17.4 percent and marginal income tax rates increase from 13.3 to 23.7 and 32.0 percent. We also
match the current pension contribution rate in Germany, but pension benefits are too high in the model. Note that due to the low fertility rate we end up with a negative native population growth rate of 1.5 percent. In the initial equilibrium we assume that immigration completely neutralizes this effect so that total population growth is zero.

Table 3 reports the participation rates of women in the model and in the German labor force. We assume that women are participating in the labor market when they work more than five percent of their time endowment. Given this definition, our model replicates the situation of women in the German labor market quite well. First, we match the average participation rate of 69.7 percent almost perfectly. Second, as in reality, participation rates increase with skill level. Third, the model also yields a close approximation of the life-cycle behavior of female labor supply which increase in the years when children attend school and decreases sharply before retirement. This pattern is similar in all skill classes but the profile is very steep in the

Table 3: Female labor market participation rates (FLMP)

<table>
<thead>
<tr>
<th></th>
<th>20-33</th>
<th>34-53</th>
<th>54-64</th>
<th>Total</th>
<th>Germany 2007\textsuperscript{a}</th>
</tr>
</thead>
<tbody>
<tr>
<td>Low-skill ($s = 1$)</td>
<td>54.1</td>
<td>64.1</td>
<td>26.9</td>
<td>51.4</td>
<td>56.7</td>
</tr>
<tr>
<td>Medium-skill ($s = 2$)</td>
<td>67.8</td>
<td>87.0</td>
<td>58.4</td>
<td>73.7</td>
<td>75.1</td>
</tr>
<tr>
<td>High-skill ($s = 3$)</td>
<td>76.5</td>
<td>96.5</td>
<td>78.1</td>
<td>85.6</td>
<td>84.4</td>
</tr>
<tr>
<td>Total</td>
<td>65.9</td>
<td>82.8</td>
<td>53.9</td>
<td>70.5</td>
<td>69.7</td>
</tr>
<tr>
<td>Germany 2006\textsuperscript{b}</td>
<td>61.0</td>
<td>73.0</td>
<td>45.0</td>
<td>–</td>
<td>–</td>
</tr>
</tbody>
</table>

Source: \textsuperscript{a}OECD(2009c), \textsuperscript{b}StaBu (2009c).
low-skill class while it is fairly flat in the high-skill class. Of course, this reflects the differences in the numbers of children.

5 Simulation results

This section presents our simulation results for the small open economy. The first subsection discusses alternative reforms of the family benefit structure whereas the second subsection concentrates on reforms of family taxation. In order to quantify the impact of a changing family policy on macroeconomic variables and welfare, we compute a new long-run equilibrium after the introduction of alternative policy reforms and compare it to the initial equilibrium discussed in Tables 2 and 3. In all simulations we assume a constant general government expenditure per capita, a constant debt-to-output ratio and balance the public budget by adjusting the consumption tax rate.

In order to separate first-order fiscal consequences from effects of changes in the population structure, we split each reform simulation in two scenarios. In the first scenario we adjust the immigration rate $\bar{\eta}$ in order to keep the aggregate growth rate of the economy constant. Then we simulate the same reform with an unaltered growth rate from immigration. In this scenario a change in native fertility affects the aggregate growth rate of the economy so that the long-run population structure changes. The latter has a direct impact on pay-as-you-go financed social security but also on the structure of public consumption and tax revenues.

5.1 Reform scenarios for family benefits

With respect to direct and indirect family benefits we consider three different policy reforms. First we increase the direct monetary transfers $b_c$ by roughly 25 percent. This could be motivated by the recent increase in child benefits and the introduction of the parental benefit (“Elterngeld”) in Germany. We compare this reform with an alternative policy that increases the per capita outlays for in-kind benefits by the same amount. If public care facilities for little children and pupils increase, children stay less at home so that time costs for their mothers decrease. Consequently, we assume in this scenario that the time costs for a pre-school child fall from 8 to 4 percent and the time costs for a pupil decrease from 5 to 2.5 percent of the available time. While both initial reforms imply an increase of per capita transfers to families, the third experiment keeps the aggregate benefits per child constant but changes their structure. In this case the increase of in-kind benefits from the second policy reform is combined

\footnote{We have also simulated the reforms in a closed economy, but the resulting repercussion effects from changing factor prices are not significant. Simulation results are available upon request.}
with an equivalent reduction in direct transfers per child. Table 4 reports the changes of some central variables.

Table 4: Fertility, macro and welfare effects of child and in-kind benefit reforms

<table>
<thead>
<tr>
<th></th>
<th>Child benefit increase</th>
<th>In-kind benefit increase</th>
<th>Benefit structure reform</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>exogenous growth</td>
<td>endogenous growth</td>
<td>exogenous growth</td>
</tr>
<tr>
<td>Fertility rates</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>TFR</td>
<td>1.82</td>
<td>1.79</td>
<td>1.92</td>
</tr>
<tr>
<td>TFR(1)</td>
<td>2.58</td>
<td>2.50 (28)</td>
<td>2.54</td>
</tr>
<tr>
<td>TFR(2)</td>
<td>1.66</td>
<td>1.62 (21)</td>
<td>1.77</td>
</tr>
<tr>
<td>TFR(3)</td>
<td>1.24</td>
<td>1.26 (14)</td>
<td>1.47</td>
</tr>
<tr>
<td>Macroeconomic effects$^a$</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$\eta + \bar{\eta}$</td>
<td>0.0</td>
<td>0.7</td>
<td>0.0</td>
</tr>
<tr>
<td>$b^\gamma M/Y$</td>
<td>1.8</td>
<td>1.8</td>
<td>1.1</td>
</tr>
<tr>
<td>$G_c/Y$</td>
<td>0.5</td>
<td>0.9</td>
<td>1.9</td>
</tr>
<tr>
<td>$\tau_c$</td>
<td>4.1</td>
<td>3.1</td>
<td>5.3</td>
</tr>
<tr>
<td>$\tau$</td>
<td>0.0</td>
<td>-3.9</td>
<td>0.0</td>
</tr>
<tr>
<td>$A_{(\text{p.c.})}$</td>
<td>3.0</td>
<td>0.1</td>
<td>0.9</td>
</tr>
<tr>
<td>$L^j_{(\text{p.c.})}$</td>
<td>-10.8</td>
<td>-5.0</td>
<td>-8.9</td>
</tr>
<tr>
<td>$L^m_{(\text{p.c.})}$</td>
<td>0.0</td>
<td>3.3</td>
<td>0.0</td>
</tr>
<tr>
<td>$Y, L_{(\text{p.c.})}$</td>
<td>-3.1</td>
<td>0.9</td>
<td>-2.6</td>
</tr>
<tr>
<td>Welfare effects$^b$</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$W(1)$</td>
<td>-0.63</td>
<td>0.99</td>
<td>-1.32</td>
</tr>
<tr>
<td>$W(2)$</td>
<td>-1.73</td>
<td>-0.05</td>
<td>-2.00</td>
</tr>
<tr>
<td>$W(3)$</td>
<td>-1.82</td>
<td>-0.15</td>
<td>-2.14</td>
</tr>
</tbody>
</table>

$^a$Change in percentage points. $^b$As a percentage of the present value of remaining resources. p.c. per capita

When we increase child benefits in the left part of Table 4, the total fertility rate rises significantly from 1.46 to 1.82. However, this increase is quite different in the three skill classes. As one would expect, child benefits have a very strong impact on low-skilled households whose total fertility rate TFR(1) increases from 1.96 to 2.58 children. At the same time, high-skilled families increase fertility only slightly from 1.11 to 1.24 children. In the long run child benefits have increased from 3.0 to 4.8 percent of GDP. About half of this increase is due to the direct effect of higher benefits (with a constant number of children) and the other half is due to the increase in the number of children. Of course, when the number of children rises, in-kind benefits have to increase as well. Due to higher children outlays the consumption tax rate has to increase by 4.1 percentage points. However, social security contributions remain constant, since the population structure is not altered. Implicitly, the reform transfers resources from old
age, when consumption is high, towards younger ages when households have children. Therefore, savings per capita increase which (at least) partly neutralizes government policy. Due to higher fertility, female labor supply falls by almost 11 percentage points. Of course, the fall in participation rates depends on age and skill level. While participation rates of young women in the low-skill class fall from 54 to 44 percent, they decrease for elderly from the top-skill class only from 78 to 77 percent. The reduction in female employment induces an outflow of capital so that the output decreases by about 3 percent. Not surprisingly, higher taxes and the reduction of per capita output induce a welfare loss for all households. Note, however, that high-skill households are hurt the most due to their lower fertility.

In the next column we keep the immigration growth rate constant so that the population growth rate increases after the reform. Consequently, the long-run dependency ratio decreases from 36.4 to 29.3 percent which in turn reduces the social security contribution rate by 3.9 percentage points. The change in the population structure reduces the consumption share of GDP and increases the investment share. Nevertheless, the consumption tax rate falls compared to the previous simulation for two reasons. First, now the government has to run a deficit in order to keep the debt-to-output ratio fixed. Second, the share of general public consumption in GDP falls since output per capita increases while the government consumption per capita remains constant. The output per capita increases due to higher male employment per capita. Lower tax and contribution rates have a positive impact on welfare. Now the low-skilled class even gains significantly, while middle-skilled and high-skilled classes lose slightly.

The middle part of Table 4 reports the consequences of an equivalent increase in in-kind benefits which reduces child-related costs for mothers. Consequently, in-kind benefits relative to GDP now increase by 1.9 percentage points while direct monetary transfers to households only increase by 1.1 percentage points due to the indirect effect. In our calibration, this reform increases aggregate fertility rates even slightly stronger compared to the previous simulation. However, in contrast to the previous reform, now the fertility rate of the high-skilled class increases even stronger than that of the low skilled class. Note that now female labor supply falls less than before, although fertility is higher. Participation rates of women are now significantly higher than in the respective previous simulation. Especially in their medium years, when the children attend schools, women work more than before. Due to higher fertility, consumption taxes have to increase more than in the respective previous simulation so that welfare losses are higher in all skill-classes.

When we allow for the repercussion of the growth rate in the forth column of Table 4, the aggregate growth rate now amounts to 0.9 percent. Again the increase in consumption taxes is more than balanced by the fall in social security contribution rates. The labor supply now

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6 Male labor supply is fixed but the change in the population structure increases per capita male employment.
increases much stronger so that the output per capita rises by 2.4 percentage points. Again, low-skilled families benefit from such a reform while high-skilled slightly lose in the long run.

In the right part of Table 4 we combine an increase of in-kind benefits with an identical reduction of direct child benefits. As a consequence, child benefits decrease now by roughly 0.5 percentage points while in-kind benefits increase from 3 percent to 3.8 percent of GDP. Not surprisingly, the rise in fertility is only modest, since now especially high-skilled women increase childbirth while low-skilled families even reduce childbirth compared to the benchmark situation in Table 2. Therefore, aggregate population growth in the last column only rises by 0.1 percentage points. Since changes in tax and contribution rates are only small, government policy now mainly transfers resources within younger cohorts. As a consequence of the increased available time, female labor supply rises in both scenarios quite significantly by 1.7 and 2.9 percent. Female participation rates increase in all skill classes but especially for low-skilled women who have less children. The higher employment reduces aggregate savings so that capital inflows increase while long-run output per capita rises by 0.5 and 1.3 percent. Therefore, similar as Da Rocha and Fuster (2006) our model also allows for a joint increase in fertility and female employment. However, as one can see in the lower part of Table 4, such a policy is not without distributional cost. Mainly middle- and high-skilled families benefit from such a reform while low-skilled households are slightly hurt.

5.2 Reform scenarios for family taxation

The existing system of taxing married couples jointly in Germany has been under critique for quite some time. On the one side it has very negative incentive effects for the second earner, since marginal income tax rates are identical for both partners in the marriage. Some experts even argue that the German income tax system is mainly responsible for the low labor market participation rate of married women. On the other side, the system has very negative distributional consequences since children are not taken into account and the tax reduction from splitting for a one-earner family rises with the income level. Since the German Basic Constitutional Law does not allow a reform towards individual taxation, various reform proposals have been put forward recently which favor some form of “family tax splitting” such as practiced in France. The general income tax function then changes to

\[ T(y_j) = 1.052 \times \Gamma(D_j) \times T05(y_j/\Gamma(D_j)), \]

where the function \( \Gamma(D_j) \) computes the splitting factor given a family with demographic characteristics \( D_j \).

The French income tax system consists of splitting factors for each spouse equal 1 (as in
Germany) and additional splitting factors of 0.5 per child for the first and the second child and 1 for each additional child. Since such a reform is quite expensive\textsuperscript{7}, we contrast it with two alternatives which would reduce income tax revenues much less. The scenario “Childless taxation” reduces the splitting factor of each spouse to 0.75 while all children receive the splitting factor 0.5. As a consequence, families without children have to pay higher income taxes than in the benchmark, families with one child are not directly affected and families with two or more children pay less taxes. In the third scenario “Large family subsidy” we keep the splitting factor of both spouses at 1.0 but introduce an additional factor of 1.0 starting with the second child. Consequently, families without children and with only one child are not affected and families with two or more children pay less taxes. Table 5 reports the simulation results of the experiments with alternative family taxation reforms.

<table>
<thead>
<tr>
<th>Table 5: Fertility, macro and welfare effects of family taxation reforms</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
</tr>
<tr>
<td></td>
</tr>
<tr>
<td>Fertility rates</td>
</tr>
<tr>
<td>TFR</td>
</tr>
<tr>
<td>TFR(1)</td>
</tr>
<tr>
<td>TFR(2)</td>
</tr>
<tr>
<td>TFR(3)</td>
</tr>
<tr>
<td>Macroeconomic effects\textsuperscript{a}</td>
</tr>
<tr>
<td>(\eta + \bar{\eta})</td>
</tr>
<tr>
<td>(b^{*}M/Y)</td>
</tr>
<tr>
<td>(G_c/Y)</td>
</tr>
<tr>
<td>(\tau_c)</td>
</tr>
<tr>
<td>(\tau)</td>
</tr>
<tr>
<td>(A) (p.c.)</td>
</tr>
<tr>
<td>(L_f) (p.c.)</td>
</tr>
<tr>
<td>(L^m) (p.c.)</td>
</tr>
<tr>
<td>(Y, L) (p.c.)</td>
</tr>
<tr>
<td>Welfare effects\textsuperscript{b}</td>
</tr>
<tr>
<td>(W(1))</td>
</tr>
<tr>
<td>(W(2))</td>
</tr>
<tr>
<td>(W(3))</td>
</tr>
</tbody>
</table>

\textsuperscript{a}Change in percentage points. \textsuperscript{b}As a percentage of the present value of remaining resources.

p.c. per capita

\textsuperscript{7}In France, the tax gains from the child splitting factors are limited to 2100 € per year. But this is not included in our simulation.
The introduction of child splitting factors similar as in the French income tax system reduces income tax revenues significantly so that the consumption tax rate has to be increased by 3.6 percentage points. However, such a tax system also generates higher fertility. Since income tax reductions rise with income, especially high-skilled families increase the number of children. Higher fertility rates now increase direct child benefits and in-kind benefits by the same amount. Compared to the previous subsection, aggregate savings now increase much stronger since mainly high-skilled families benefit from the redistribution towards younger cohorts. As before, aggregate employment and output fall due to the reduction in female labor supply. Surprisingly, families from the high-skilled class are hurt the most despite the fact that they benefit the most per child. Since this class has the highest fraction of childless families, their total (average) tax benefit might be still smaller than the one for low-skilled families. The latter even gain, when we include all repercussions of the higher fertility rate for the whole economy. In this case the social security contribution rate decreases by 3.6 percentage points and per capita output rises slightly. For high-skilled families the negative effects from higher taxes still dominate so that they experience a welfare loss even in the simulation with endogenous growth.

As already mentioned before, high-skilled families have the highest fraction of childless. For this reason they lose the most in terms of welfare when we reduce the splitting factor for both spouses from 1 to 0.75. As before, the reformed tax system increases fertility rates significantly and especially for high-skilled families. Consequently, direct and in-kind benefits for families with children increase now by 1.3 and 1.5 percentage points. Of course, part of this increase in outlays is financed by higher income tax revenues from families either childless or where the children have already left the household. Consequently, the consumption tax rate now only rises by 2.2 and 1.3 percentage points. On the other hand, female employment now falls dramatically by 18.5 and 10.3 percent since marginal tax rates for families without children increase significantly. Despite the fact that output per capita increases in the endogenous growth scenario, all families are worse off in the long run. Again, high-skilled families are hurt the most since they face the strongest increase in marginal tax rates.

In the last simulation of Table 5 we keep the splitting factor of both spouses unaltered and introduce an additional splitting factor starting with the second child. Consequently, families without children and those with only one child are not directly affected by the reform. Since tax savings rise with income, fertility mainly rises for medium-skilled and high-skilled families. The fiscal redistribution is smaller then in the previous scenarios, which in turn dampens the impact on birth rates. Again, due to the intergenerational redistribution towards younger households aggregate savings rise. At the same time the reduction of female labor supply is dampened so that output now increases the most in the endogenous growth scenario. Within cohorts welfare effects are now very evenly distributed so that all family types gain significantly when the repercussions from the changes in the population structure are taken into account.
6 Conclusion

Summing up the results from the previous section, we have shown that our model is able to analyze a broad range of policy reforms that are implemented to increase fertility rates in developed economies. In all simulation experiments considered, fertility rates increase but with different intensities within and across the cohorts. Higher direct or indirect family benefits financed by consumption taxation mainly increase fertility and welfare of low-skilled families while high-skilled families are hurt in the long run. As one would expect from the traditional household theory, higher fertility rates also reduce female labor supply in this case. But it is also possible to combine an increase in fertility and female employment if the structure of family benefits changes from direct towards in-kind benefits. However, this policy has negative distributional implications since especially low-skilled families are the long-run losers.

The results from alternative forms of family taxation also indicate the effectiveness of this policy instrument with respect to the fertility rate. However, such policies are much likely to reduce female labor force participation strongly and decrease the long-run welfare of especially high-skilled families. It turns out that a policy that increases the splitting factor for families with two or more children has the most beneficial consequences. In this case, the increase in birth rates is financed by a modest increase in consumption taxation so that female labor supply only falls slightly. As a consequence, welfare of all skill classes increases in the long run.

Of course, one has to be careful to take the model too serious and draw too many policy conclusions from the reported results. At the moment our model structure is still very basic and abstracts from many real world features that are very important for fertility and female labor supply decisions. The most important deficiency seems to be the fact that our model only considers uncertainty with respect to childbirth whereas employment is always certain. In the future we will integrate especially the impact of labor market frictions and the uncertainty of future employment opportunities. Similar as Da Rocha and Fuster (2006) we plan to model a labor market where woman search for a job and working females lose jobs with an exogenous probability. In addition, we will further disaggregate within a cohort in order to distinguish between singles (who are always without children) and families (with or without children). Finally, we will include a transition period to the long-run equilibrium in order to isolate the intergenerational welfare consequences of policy reforms.
References


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